MA4270 Data Modelling and Computation

Preliminaries

- Variance = $\mathbb{E}[(X \mathbb{E}[X])^2] = \mathbb{E}[X^2] \mathbb{E}[X]^2$
- Markov ineq.: Given a > 0 and X nonnegative, $\Pr(X \ge a) \le \frac{\mathbb{E}[X]}{a}$
- Chebyshev ineq.: $\Pr\left(|Y \mathbb{E}[Y]| \ge b\right) \le \frac{\sigma_Y^2}{b^2}$

Binary Classification

- Training data: $\mathcal{D} = \{(\mathbf{x}_t, y_t)\}_{t=1}^n$ - where $y_t \in \{-1, 1\}$ and $\mathbf{x}_t \in \mathbb{R}^d$ - where we have *n* training data
- Classifier: $f_{\theta} : \mathbb{R}^d \to \{-1, 1\}$
- Training error: $\hat{E}(\mathbf{\theta}) = \frac{1}{n} \sum_{t=1}^{n} \text{Loss}(y_t, f_{\mathbf{\theta}}(\mathbf{x}_t))$ - i.e. the proportion of failed training data
- Loss $(y, \hat{y}) = \mathbf{1}\{\hat{y} \neq y\}$

Linear classifier

- Formula: ŷ := sgn⟨θ, x⟩ := sgn ∑_{i=1}^d θ_ix_i (for some fixed θ
 i.e. partitioning space with plane passing through origin
- \mathcal{D} is linearly separable: $\exists \boldsymbol{\theta}$ such that $\hat{E}(\boldsymbol{\theta}) = 0$

• Perceptron update algorithm:

- current estimate θ_{curr}
- look at one (\mathbf{x}_t, y_t)
- if $\operatorname{sgn}\langle \boldsymbol{\theta}_{curr}, \mathbf{x}_t \rangle = y_t$ then $\boldsymbol{\theta}_{next} \coloneqq \boldsymbol{\theta}_{curr}$

- otherwise set $\boldsymbol{\theta}_{next} \coloneqq \boldsymbol{\theta}_{curr} + y_t \mathbf{x}_t$

(intuitively, we are increasing the value of $y_t \boldsymbol{\theta}_{next}^T \mathbf{x}_t$)

• Perceptron algorithm:

- init $\mathbf{\theta}^{(0)} \coloneqq \mathbf{0}, k \coloneqq 0$

For each data point, run the update, and increment k only when we made a mistake (repeat after reaching n)
stop when we do full pass of all data without mistakes

• Perceptron algorithm assumptions:

- $\exists R > 0$ such that $\forall t, ||\mathbf{x}_t|| \leq R$ (i.e. \mathbf{x}_t is uniformly bdd) - $\exists \boldsymbol{\theta}^*$ and $\exists \gamma > 0$ such that $\min_t y_t(\boldsymbol{\theta}^*)^T \mathbf{x}_t \geq \gamma$ (i.e. stronger lin. separable cond. – bounded away from zero) **Thm**: under above assumptions, the algorithm finds $\boldsymbol{\theta}^{(k)}$ such that $\hat{E}(\boldsymbol{\theta}^{(k)}) = 0$ after at most $k_{max} = \frac{R^2 ||\boldsymbol{\theta}^*||^2}{\gamma^2}$ steps (we make at most k_{max} mistakes)

- Cauchy-Schwarz inequality: $|\langle \mathbf{u}, \mathbf{v} \rangle| \le ||\mathbf{u}|| \cdot ||\mathbf{v}||$ (with equality when \mathbf{u} and \mathbf{v} are in the same direction)
- Perceptron proof:

1. show that $\langle \boldsymbol{\theta}^*, \boldsymbol{\theta}^{(k+1)} \rangle \geq \langle \boldsymbol{\theta}^*, \boldsymbol{\theta}^{(k)} \rangle + \gamma$, so $\langle \boldsymbol{\theta}^*, \boldsymbol{\theta}^{(k)} \rangle \geq k\gamma$ 2. show that $\left\| \boldsymbol{\theta}^{(k+1)} \right\|^2 \leq \left\| \boldsymbol{\theta}^{(k)} \right\|^2 + R^2$, so $\left\| \boldsymbol{\theta}^{(k)} \right\|^2 \leq kR^2$ 3. use Cauchy-Schwartz ineq. to conclude that $k \leq \frac{R^2 \|\boldsymbol{\theta}^*\|^2}{\gamma^2}$ Details: 1. $(\boldsymbol{\theta}^*)^T \boldsymbol{\theta}^{(k+1)} = (\boldsymbol{\theta}^*)^T (\boldsymbol{\theta}^{(k)} + y_t \mathbf{x}_t) =$

1. (**b**') $^{T} \mathbf{\theta}^{(k)} + y_t(\mathbf{\theta}^*)^T \mathbf{x}_t \ge (\mathbf{\theta}^*)^T \mathbf{\theta}^{(k)} + \gamma$ 2. $\|\mathbf{\theta}^{(k+1)}\|^2 = \|\mathbf{\theta}^{(k)} + y_t \mathbf{x}_t\|^2 =$ $\|\mathbf{\theta}^{(k)}\|^2 + 2\langle \mathbf{\theta}^{(k)}, y_t \mathbf{x}_t \rangle + \|\mathbf{x}_t\|^2 \le \|\mathbf{\theta}^{(k)}\|^2 + \|\mathbf{x}_t\|^2$ (since

when a mistake occurs we must have $\langle \boldsymbol{\theta}^{(k)}, y_t \mathbf{x}_t \rangle \leq 0$

3.
$$1 \ge \frac{\langle \boldsymbol{\theta}^*, \boldsymbol{\theta}^{(k)} \rangle}{\|\boldsymbol{\theta}^*\| \cdot \|\boldsymbol{\theta}^{(k)}\|} \ge \frac{k\gamma}{\|\boldsymbol{\theta}^*\| \cdot \sqrt{kR^2}} = \frac{\sqrt{k\gamma}}{\|\boldsymbol{\theta}^*\| \cdot R}, \text{ so } k \le \frac{R^2 \|\boldsymbol{\theta}^*\|^2}{\gamma^2}$$

Perceptron with offset:
Set $\hat{\boldsymbol{\theta}} \coloneqq \begin{bmatrix} \boldsymbol{\theta} \\ \theta_0 \end{bmatrix}$ and $\hat{\mathbf{x}} \coloneqq \begin{bmatrix} \mathbf{x} \\ 1 \end{bmatrix}$,
so that $\hat{\boldsymbol{\theta}}^T \hat{\mathbf{x}} = \boldsymbol{\theta}^T \mathbf{x} + \theta_0$ and $\hat{R}^2 = R^2 + c^2$

- (Non-geometric) margin (w.r.t. $\mathbf{\theta}^*$): $\gamma \coloneqq \min_t y_t(\mathbf{\theta}^*)^T \mathbf{x}_t$ (note: usually "margin" means the geometric margin)
- Geometric margin (normalised by $\|\boldsymbol{\theta}^*\|$): $\gamma_{geom} \coloneqq \frac{\gamma}{\|\boldsymbol{\theta}^*\|}$ (= the shortest distance from a point to the hyperplane)
- Maximum margin classifier (a kind of support vector machine): find the classifier that makes the margin largest
 maximise (over θ, γ) γ/||θ|| subject to ∀t, ytθ^Txt ≥ γ
 equiv. minimise (over new θ) 1/2 ||θ||² subject to
- $\forall t, y_t \boldsymbol{\theta}^T \mathbf{x}_t > 1$

.

- which is a convex optimisation problem solvable efficiently
- then the classifier $f_{\theta}(\mathbf{x}) = \operatorname{sgn}\left(\boldsymbol{\theta}^T \mathbf{x}\right)$ is max margin - and $\gamma_{geom} = \frac{1}{\|\boldsymbol{\theta}^*\|}$ (new $\boldsymbol{\theta}$)

Proof of uniqueness: can show that if we have distinct θ_1 , θ_2 both optimal, then their average most also be optimal and hence $\|\theta_1 - \theta_2\|^2 = 0$

• Support vector: the xs that touch the margin

Linear classifier with offset

- Formula: $\hat{y} \coloneqq \operatorname{sgn}(\langle \boldsymbol{\theta}, \mathbf{x} \rangle + \theta_0)$ where $\boldsymbol{\theta} \in \mathbb{R}^d$ and $\theta_0 \in \mathbb{R}$
- SVM with offset: minimise (over $\boldsymbol{\theta}, \theta_0$) $\frac{1}{2} \|\boldsymbol{\theta}\|^2$ subject to $\forall t, y_t \left(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0\right) \ge 1$ (note: we don't penalise for θ_0)
- Soft-margin SVM: allows misclassified points by introducing slack variables (ζ):

- minimise (over $\mathbf{\theta}$, θ_0 , ζ) $\frac{1}{2} \|\mathbf{\theta}\|^2 + C \sum_{t=1}^n \zeta_t$ subject to

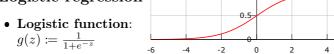
 $\forall t, y_t \left(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0 \right) \geq 1 - \zeta_t \text{ and } \forall t, \zeta_t \geq 0 \text{ (allows us to pay a measured penalty for points too close to the margin or misclassified)}$

- equiv. minimise (over $\boldsymbol{\theta}, \theta_0$) $\frac{1}{2} \|\boldsymbol{\theta}\|^2 + C \sum_{t=1}^n \left[1 - y_t \left(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0 \right) \right]_+$ where $[z]_+ := \max\{0, z\}$

- higher $C \implies$ favour fewer violations

- $C \to \infty$ and \mathcal{D} is lin. separable \implies hard-margin SVM **Support vector**: the **x**s that touch the margin, are inside the margin, or misclassified
- moving these points will change the line
- if we remove all points except those from support vector, the classifier will remain the same
- Loss: (where $z = y_t \left(\boldsymbol{\theta}^T \mathbf{x}_t + \boldsymbol{\theta}_0 \right)$) - $\frac{0 - 1 \text{ loss: } Loss(z) \coloneqq \mathbf{1} \{ z \le 0 \}$
 - <u>Hinge loss</u>: $Loss_h(z) \coloneqq [1-z]_+$ (it is convex)

Logistic regression



• Soft decision: We answer using a probability space in [0, 1] (logistic regression is a soft-decision algorithm)

• **Probability** (given
$$\boldsymbol{\theta}$$
, θ_0):
 $P(y = 1 \mid \mathbf{x}) = g\left(\boldsymbol{\theta}^T \mathbf{x} + \theta_0\right) = \frac{1}{1 + \exp\left(-\left(\boldsymbol{\theta}^T \mathbf{x} + \theta_0\right)\right)}$
 $P(y = -1 \mid \mathbf{x}) = 1 - P(y = 1 \mid \mathbf{x}) = \frac{1}{1 + \exp\left(\boldsymbol{\theta}^T \mathbf{x} + \theta_0\right)}$
Hence $P(y \mid \mathbf{x}) = g\left(y \cdot \left(\boldsymbol{\theta}^T \mathbf{x} + \theta_0\right)\right)$
Note: $\log\left(\frac{P(y=1|\mathbf{x})}{P(y=-1|\mathbf{x})}\right) = \boldsymbol{\theta}^T \mathbf{x} + \theta_0$

- Multiplying θ and θ_0 by a large constant makes the probability further away from 0.5 (i.e. the graph gets steeper, but the decision boundary remains the same)
- To determine θ and θ_0 , we generally use "maximum likelihood": $L(\theta, \theta_0 \mid \mathcal{D}) = \prod_{t=1}^n P(y_t \mid \mathbf{x}_t; \theta, \theta_0)$ (generally good enough if *n* is large) Equiv. $L(\theta, \theta_0 \mid \mathcal{D}) = P(y_1, \dots, y_n \mid \mathbf{x}_1, \dots, \mathbf{x}_n; \theta, \theta_0)$ max likelihood: $(\theta, \theta_0) = \arg \max_{\theta, \theta_0} L(\theta, \theta_0 \mid \mathcal{D})$ - To solve: $\arg \max_{\theta, \theta_0} L(\theta, \theta_0 \mid \mathcal{D})$ = $\arg \max_{\theta, \theta_0} \prod_{t=1}^n P(y_t \mid \mathbf{x}_t; \theta, \theta_0)$ = $\arg \max_{\theta, \theta_0} \sum_{t=1}^n \log P(y_t \mid \mathbf{x}_t; \theta, \theta_0)$ = $\arg \min_{\theta, \theta_0} \sum_{t=1}^n \log(1 + \exp(-y_t(\theta^T \mathbf{x}_t + \theta_0))))$... which can only be solved numerically using gradient descent, because it is <u>convex</u> - We can call it a "loss": <u>Logistic loss</u>: $Loss(z) = Log(1 + e^{-z})$
- Gradient descent for max likelihood logistic regression: $\frac{\partial}{\partial \theta_0} \operatorname{Log}(1 + \exp(-y_t(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0))) = -y_t (1 - P(y_t \mid \mathbf{x}_t; \boldsymbol{\theta}, \theta_0))$ $\frac{\partial}{\partial \boldsymbol{\theta}} \operatorname{Log}(1 + \exp(-y_t(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0))) = -y_t \mathbf{x}_t (1 - P(y_t \mid \mathbf{x}_t; \boldsymbol{\theta}, \theta_0))$ - Thus: $\theta_0^{(i+1)} \leftarrow \theta_0^{(i)} + \eta \sum_{i=1}^n y_t \left(1 - P\left(y_t \mid \mathbf{x}_t; \boldsymbol{\theta}^{(n)}, \theta_0^{(n)}\right)\right)$ $\boldsymbol{\theta}^{(i+1)} \leftarrow \boldsymbol{\theta}^{(i)} + \eta \sum_{i=1}^n y_t \mathbf{x}_t \left(1 - P\left(y_t \mid \mathbf{x}_t; \boldsymbol{\theta}^{(n)}, \theta_0^{(n)}\right)\right)$
- Stochastic gradient descent:

- For large data sets, iterating all points at every step is too costly, so we instead pick one (or a few) data points only (at random, or cycle, etc):

$$\theta_0^{(i+1)} \leftarrow \theta_0^{(i)} + \eta y_t \left(1 - P\left(y_t \mid \mathbf{x}_t; \boldsymbol{\theta}^{(n)}, \theta_0^{(n)}\right) \right) \\ \boldsymbol{\theta}^{(i+1)} \leftarrow \boldsymbol{\theta}^{(i)} + \eta y_t \mathbf{x}_t \left(1 - P\left(y_t \mid \mathbf{x}_t; \boldsymbol{\theta}^{(n)}, \theta_0^{(n)}\right) \right)$$

- For separable data, max likelihood will lead to an arbitrarily steep logistic regression function (i.e. $L(\theta, \theta_0 \mid D) \rightarrow 1 \text{ as } \theta, \theta_0 \rightarrow \infty$ is the max likelihood), so we want to prevent this using <u>regularisation</u>
- **Regularisation**: Change the expression to $\arg\min_{\boldsymbol{\theta},\theta_0} \sum_{t=1}^n \text{Log}(1 + \exp(-y_t(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0))) + \frac{\lambda}{2} \|\boldsymbol{\theta}\|^2$ (where λ is the regularisation parameter) (prevents overly large $\boldsymbol{\theta}$)
- Logistic regression can be extended to multi-class classification

Gradient descent

- Def: Want to find $\hat{\mathbf{z}} = \arg\min_{\mathbf{z} \in \mathbb{R}^d} f(\mathbf{z})$ for some evaluation function f
- Algorithm: At the current point, evaluate the current gradient, and move a bit in the steepest downhill direction
 update operation: z⁽ⁱ⁺¹⁾ ← z⁽ⁱ⁾ − η∇f(z⁽ⁱ⁾)

where $\nabla f(\mathbf{z}^{(i)}) \coloneqq \begin{bmatrix} \frac{\partial f}{\partial z_1} \\ \vdots \\ \frac{\partial f}{\partial x_1} \end{bmatrix}$ and η is the step size

- guaranteed to find the minimum if the function is convex

Linear regression

• The output is now over \mathbb{R}

• Training data: $\mathcal{D} = \{(\mathbf{x}_t, y_t)\}_{t=1}^n$ - where $y_t \in \mathbb{R}$ and $\mathbf{x}_t \in \mathbb{R}^d$

• **Objective**: To learn a prediction rule (line): $\hat{y} = \theta^T \mathbf{x} + \theta_0$

• Gaussian model: Add some noise, because the relationship may not be perfect - $y = (\mathbf{\theta}^*)^T \mathbf{x} + \theta_0^* + z$, where $z \sim N(0, \sigma^2)$ - PDF of a normal distribution: $\mathcal{N}(z; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(\frac{-(z-\mu)^2}{2\sigma^2}\right)$ - So $P(y \mid \mathbf{x}) = \mathcal{N}\left(y; (\mathbf{\theta}^*)^T \mathbf{x} + \theta_0^*, \sigma^2\right)$

• Using maximum likelihood to derive least squares formula (works for Gaussian noise model only): $L(\boldsymbol{\theta}, \theta_0, \sigma^2 \mid \mathcal{D}) = P(y_1, \dots, y_n \mid \mathbf{x}_1, \dots, \mathbf{x}_n; \boldsymbol{\theta}, \theta_0, \sigma^2) = \prod_{t=1}^n P(y_t \mid \mathbf{x}_t; \boldsymbol{\theta}, \theta_0, \sigma^2) = \prod_{t=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_t - \boldsymbol{\theta}^T \mathbf{x}_t - \theta_0)^2}{2\sigma^2}\right) \\ \text{Log}(L(\boldsymbol{\theta}, \theta_0, \sigma^2 \mid \mathcal{D})) = -\frac{n}{2} \text{Log}(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{t=1}^n \left(y_t - \boldsymbol{\theta}^T \mathbf{x}_t - \theta_0\right)^2 \\ \dots \text{ so } \boldsymbol{\theta} \text{ and } \theta_0 \text{ do not depend on } \sigma^2 \\ \dots \text{ so equiv. to finding} \\ \left(\hat{\boldsymbol{\theta}}, \hat{\theta}_0\right) = \arg\min_{\boldsymbol{\theta}, \theta_0} \sum_{t=1}^n \left(y_t - \left(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0\right)\right)^2 \\ (\text{note that } \left(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0\right) \text{ is simply the predicted value of } y \\ \text{ using the line defined by } \left(\hat{\boldsymbol{\theta}}, \hat{\theta}_0\right), \text{ so it is the formula for least squares regression} \end{cases}$

• Closed form solution (for the prediction/estimate): Solve $J(\boldsymbol{\Theta}) \coloneqq \sum_{t=1}^{n} \left(y_t - \left(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0 \right) \right)^2 = \| \mathbf{y} - \mathbf{X} \boldsymbol{\Theta} \|^2$, where: $- \mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \in \mathbb{R}^n$; $\mathbf{X} = \begin{bmatrix} \mathbf{x}_1^{T-1} \\ \vdots \\ \mathbf{x}_n^{T-1} \end{bmatrix} \in \mathbb{R}^{n \times (d+1)}$; $\boldsymbol{\Theta} = \begin{bmatrix} \boldsymbol{\theta} \\ \theta_0 \end{bmatrix} \in \mathbb{R}^{d+1}$

then to solve for min point, we make the gradient equal 0: $\nabla \left(\|\mathbf{y} - \mathbf{X}\boldsymbol{\Theta}\|^2 \right) = -2\mathbf{X}^T \left(\mathbf{y} - \mathbf{X}\boldsymbol{\Theta} \right) = \mathbf{0}$ $\implies \mathbf{X}^T \mathbf{y} = \mathbf{X}^T \mathbf{X}\boldsymbol{\Theta}$ $\implies \boldsymbol{\Theta} = \left(\mathbf{X}^T \mathbf{X} \right)^{-1} \mathbf{X}^T \mathbf{y} \text{ (so } \boldsymbol{\Theta} \text{ is linear in } \mathbf{y} \text{ but not } \mathbf{X} \text{)}$ $\left(\left(\mathbf{X}^T \mathbf{X} \right)^{-1} \mathbf{X}^T \text{ is called the pseudo-inverse of } \mathbf{X} \text{)}$ Prediction rule: $\hat{y} = \hat{\boldsymbol{\theta}}^T \mathbf{x} + \hat{\theta}_0$

• Maximum likelihood with high *n*: low bias and low variance Maximum likelihood with low *n*: low bias and high variance

Bias & Variance

• Bias: systematic error Variance: random error

How good is our predicted Θ ?

- Goal: minimise $\mathbb{E}\left[\left\|\hat{\boldsymbol{\Theta}} \boldsymbol{\Theta}^*\right\|^2\right]$ (i.e. mean squared error) (note: $\hat{\boldsymbol{\Theta}}$ is our estimate; $\boldsymbol{\Theta}^*$ is actual)
- Bias-variance decomposition:

 $\underbrace{\mathbb{E}\left[\left\|\hat{\boldsymbol{\Theta}}-\boldsymbol{\Theta}^*\right\|^2\right]}_{\text{mean sq. error}} = \underbrace{\left\|\mathbb{E}\left[\hat{\boldsymbol{\Theta}}\right]-\boldsymbol{\Theta}^*\right\|^2}_{\text{bias squared}} + \underbrace{\mathbb{E}\left[\left\|\hat{\boldsymbol{\Theta}}-E\left[\hat{\boldsymbol{\Theta}}\right]\right\|^2\right]}_{\text{variance}}$ Note: $\mathbb{E}\left[\hat{\boldsymbol{\Theta}}\right]-\boldsymbol{\Theta}^* = -\lambda \left(\mathbf{X}^T\mathbf{X}+\lambda\mathbf{I}\right)^{-1}\boldsymbol{\Theta}^*$

- To show that least squares estimate is unbiased - Combining all data points for a true value (Θ^*): $y = (\Theta^*)^T \mathbf{x} + \theta_0^* + z$ into a vector, we get $\mathbf{y} = \mathbf{X}\Theta^* + \mathbf{z}$ - substitute equation into $\hat{\Theta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ to get $\hat{\Theta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{X}\Theta^* + \mathbf{z}) = \Theta^* + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{z}$ - and since $\mathbb{E}[\mathbf{z}] = 0$, we have $\mathbb{E}[\hat{\Theta}] = \Theta^*$
- so we conclude that $\hat{\Theta}$ is an unbiased estimate for Θ^*
- Variance term may be very high if $\mathbf{X}^T \mathbf{X}$ has small eigenvalues (intuitively, $\mathbf{X}^T \mathbf{X}$ is "almost singular")

 $\mathbb{E}\left[\left\|\hat{\boldsymbol{\Theta}} - E\left[\hat{\boldsymbol{\Theta}}\right]\right\|^{2}\right] = \operatorname{Tr}\left[\operatorname{Cov}\left[\hat{\boldsymbol{\Theta}}\right]\right], \text{ and}$ $\operatorname{Cov}\left[\hat{\boldsymbol{\Theta}}\right] = \sigma^{2} \left(\mathbf{X}^{T}\mathbf{X}\right)^{-1}, \text{ so}$ $\mathbb{E}\left[\left\|\hat{\boldsymbol{\Theta}} - E\left[\hat{\boldsymbol{\Theta}}\right]\right\|^{2}\right] = \sigma^{2} \operatorname{Tr}\left[\left(\mathbf{X}^{T}\mathbf{X}\right)^{-1}\right]$... and thus variance high when $\mathbf{X}^{T}\mathbf{X}$ has small eigenvalues

Trading bias for variance – Regularisation (with ^λ/₂ ||θ||²) (aka. "<u>ridge regression</u>"):
we want to penalize large θ (but not θ₀) (θ̂, θ̂₀)

$$= \arg\min_{\boldsymbol{\theta}, \theta_0} \sum_{t=1}^n \left(y_t - \left(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0 \right) \right)^2 + \frac{\lambda}{2} \sum_{j=1}^d \theta_j^2$$
$$= \arg\min_{\boldsymbol{\theta}, \theta_0} \| \mathbf{y} - \mathbf{X} \boldsymbol{\Theta} \|^2 + \frac{\lambda}{2} \| \boldsymbol{\theta} \|^2$$

Closed form solution: $\hat{\boldsymbol{\Theta}} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$ We want to find the best λ (perhaps try a few different λ and check performance on validation data)

Non-linear Rules

Given D := {(x_t, y_t)}ⁿ_{t=1}, we can map it to some D' := {(Φ(x_t), y_t)}ⁿ_{t=1}, and map back the found θ and θ₀
for example, if we want a quadratic expression, we can let Φ(x) = [x, x²]^T; if we want a circle, we can write an equation of a circle

- if d > 1, we can also consider cross terms, e.g. x_1x_2 terms - if there are too many features, it might be prone to overfitting

Kernel methods

- We want the new dataset to be $k(\mathbf{x}_i, \mathbf{x}_j) \coloneqq \langle \mathbf{\Phi}(\mathbf{x}_i), \mathbf{\Phi}(\mathbf{x}_j) \rangle$, where
- φ is a "feature map"
- we compute it without explicitly calculating $\mathbf{\Phi}(\mathbf{x}_i)$
- it works nice because inner products capture geometry
 this trick works only when the rule becomes linear on the

new dataset - note: the fresh \mathbf{x} whose y is being predicted usually has its kernel with existing data points calculated too, and goes

its kernel with existing data points calculated too, and goes into the prediction function

- E.g.: If $\mathbf{\Phi}(x) = (1, \sqrt{3}x, \sqrt{3}x^2, x^3)$, then we have kernel $\langle \mathbf{\Phi}(x), \mathbf{\Phi}(x') \rangle = (1 + xx')^3$ For degree *p* polynomial and *d* dimensions, we have the polynomial kernel $k(\mathbf{x}, \mathbf{x}') = (1 + \langle \mathbf{x}, \mathbf{x}' \rangle)^p$
- E.g.: String kernel, let $k(\mathbf{x}, \mathbf{x}')$ be the number of words appearing in both strings
- Intuitively (non-rigorously), $k({\bf x},{\bf x}')$ should be a kind of measure of similarity
- Def: k(x, x') is a kernel: k(x, x') = (φ(x), φ(x')) for some φ (possibly infinite-dimensional)
- Positive semidefinite: $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ is positive semidefinite: for any $\mathbf{x}, \mathbf{x}', k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$, and for any $\mathbf{x}_1, \ldots, \mathbf{x}_m$ (for any m), the Gram matrix
- $\mathbf{K} \coloneqq \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_m) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_m, \mathbf{x}_1) & \cdots & k(\mathbf{x}_m, \mathbf{x}_m) \end{bmatrix}$ is positive semidefinite

 $[\kappa(\mathbf{x}_m, \mathbf{x}_1) \quad \cdots \quad \kappa(\mathbf{x}_m, \mathbf{x}_m)]$ (note: positive semidefinite matrix \implies symmetric matrix)

- Positive semidefinite matrix: (two equiv. characterisations):
 ∀z, z^TKz ≥ 0
 all upper-left square submatrices have nonnegative determinant, or
 all eigenvalues are nonnegative
- Thm k is a kernel $\iff k$ is positive semidefinite (useful for proving if something is not a valid kernel) <u>Proof for " \implies ": let $\mathbf{K} = \mathbf{\Phi}^T \mathbf{\Phi}$ where $\overline{\mathbf{\Phi}} := [\mathbf{\Phi}(\mathbf{x}_1), \dots, \mathbf{\Phi}(\mathbf{x}_n)]$, and observe that \mathbf{K} is the Gram matrix; then observe that $\forall \mathbf{z}, \mathbf{z}^T \mathbf{K} \mathbf{z} = \mathbf{z}^T \mathbf{\Phi}^T \mathbf{\Phi} \mathbf{z} = (\mathbf{\Phi} \mathbf{z})^T (\mathbf{\Phi} \mathbf{z}) = \|\mathbf{\Phi} \mathbf{z}\|^2 \ge 0$ (so \mathbf{K} is positive semidefinite)</u>
- To show that $k(\mathbf{x}, \mathbf{x}')$ is a kernel: (any one) explicitly find $\boldsymbol{\Phi}$
- use kernel closure properties below
- (rarely, and hard) show that k is positive-semidefinite
- To show that $k(\mathbf{x}, \mathbf{x}')$ is <u>not</u> a kernel: (any one)
- k is not symmetric
- $k(\mathbf{x}, \mathbf{x}) < 0$ for some \mathbf{x}
- Find some \mathbf{x}, \mathbf{x}' such that det $\left(\begin{bmatrix} k(\mathbf{x}, \mathbf{x}) & k(\mathbf{x}, \mathbf{x}') \\ k(\mathbf{x}', \mathbf{x}) & k(\mathbf{x}', \mathbf{x}') \end{bmatrix} \right) < 0$
- **Deriving more kernels**: If k_1, k_2 are kernels, then:
- \forall functions f (real-valued), $f(\mathbf{x})k_1(\mathbf{x}, \mathbf{x}')f(\mathbf{x}')$ is a kernel - $k_1(\mathbf{x}, \mathbf{x}') + k_2(\mathbf{x}, \mathbf{x}')$ is a kernel - $k_1(\mathbf{x}, \mathbf{x}')k_2(\mathbf{x}, \mathbf{x}')$ is a kernel (proof of first: let $\mathbf{\phi}(\mathbf{x}) = \mathbf{\phi}_1(\mathbf{x})f(\mathbf{x})$ and show that $\langle \mathbf{\phi}(\mathbf{x}), \mathbf{\phi}(\mathbf{x}') \rangle = f(\mathbf{x})k_1(\mathbf{x}, \mathbf{x}')f(\mathbf{x}')$) (proof of second: let $\mathbf{\phi}(\mathbf{x}) = \begin{bmatrix} \mathbf{\phi}_1(\mathbf{x}) \\ \mathbf{\phi}_2(\mathbf{x}) \end{bmatrix}$ and show that $\langle \mathbf{\phi}(\mathbf{x}), \mathbf{\phi}(\mathbf{x}') \rangle = k_1(\mathbf{x}, \mathbf{x}') + k_2(\mathbf{x}, \mathbf{x}')$) (proof of third: let $\mathbf{\phi}(\mathbf{x})$ be a vector containing entries

(proof of third: let $\mathbf{\phi}(\mathbf{x})$ be a vector containing entries $\phi_{ij}(\mathbf{x}) = \phi_i^{(1)}(\mathbf{x})\phi_j^{(2)}(\mathbf{x})$ for each i, j, and show that $\langle \mathbf{\phi}(\mathbf{x}), \mathbf{\phi}(\mathbf{x}') \rangle = k_1(\mathbf{x}, \mathbf{x}')k_2(\mathbf{x}, \mathbf{x}'))$

• RBF kernel (aka. Gaussian kernel): $k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2} ||\mathbf{x} - \mathbf{x}'||^2)$

- measures distance between points; larger (and tends to 1) as $\mathbf{x}' \to \mathbf{x}$

- has infinite number of features $(\boldsymbol{\varphi}(\mathbf{x})),$ and it is not easy to determine

RBF kernel (aka. Gaussian kernel) with length scale: k(x, x') = exp(-¹/_{2ℓ} ||x - x'||²)
ℓ = length scale, increasing ℓ means that points don't need to be too close in order to be considered similar

Some "kernelizable" learning algorithms:
k-nearest neighbours: Want to make a prediction for x (where ŷ ∈ {-1,1}). Given x, we find the k nearest points, and predict ŷ by majority rule. This is kernelizable because ||x - x_t|| can be expressed as inner products.
Linear (ridge) regression: After solving Θ, want to predict (Θ x) given x : observe that

$$(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I}_d)^{-1} \mathbf{X}^T = \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_n)^{-1}$$
, so
 $\hat{\mathbf{\Theta}} = \mathbf{X}^T (\mathbf{X} \mathbf{X}^T + \lambda \mathbf{I}_n)^{-1} \mathbf{y}$, so

 $\hat{y} = \mathbf{x}^T \mathbf{X}^T (\mathbf{X}\mathbf{X}^T + \lambda \mathbf{I}_n)^{-1} \mathbf{y}$ (for no-offset case). Observe that $\mathbf{X}\mathbf{X}^T$ is just the Gram matrix with standard inner products (i.e. a matrix of inner products), and $\mathbf{x}^T \mathbf{X}^T$ is a vector of inner products, so we can replace them with inner products and this prediction rule (for \hat{y}) would be kernelisable, i.e. $\hat{y} = \mathbf{k}(\mathbf{x})(\mathbf{K} + \lambda \mathbf{I}_n)^{-1}\mathbf{y}$ where **K** is the

Gram matrix and $\mathbf{k}(\mathbf{x}) = \begin{bmatrix} k(\mathbf{x}, \mathbf{x}_1) \\ \vdots \\ k(\mathbf{x}, \mathbf{x}_n) \end{bmatrix}$

Convex Optimisation (for Kernel SVM)

- $\mathcal{D} \subseteq \mathbb{R}^d$ is a **convex set**: $\forall \mathbf{x}, \mathbf{x}' \in \mathcal{D}, \forall \lambda \in [0, 1], \lambda \mathbf{x} + (1 - \lambda) \mathbf{x}' \in \mathcal{D}$
- $f: D \to \mathbb{R}$ (where $D \subseteq \mathbb{R}^d$ is convex) is a **convex function**: $\forall \mathbf{x}, \mathbf{x}' \in D, \forall \lambda \in [0, 1], f(\lambda \mathbf{x} + (1 - \lambda) \mathbf{x}') \leq \lambda f(\mathbf{x}) + (1 - \lambda) f(\mathbf{x}')$ (for **concave function**, change " \leq " to " \geq ")
- Equiv. def if f differentiable: $f(\mathbf{x}') \ge f(\mathbf{x}) + \nabla f(\mathbf{x})^T (\mathbf{x}' - \mathbf{x})$ (i.e. the second point lies above the tangent from the first point)
- Equiv. def if f twice differentiable: $\nabla^2 f(\mathbf{x}) \ge \mathbf{0}$ (i.e. is positive semidefinite) (for 1D case, $f''(x) \ge 0$)
- Properties of convex functions:

- $f_1, f_2 \text{ convex} \implies$ $\forall \alpha_1, \alpha_2 \ge 0, \alpha_1 f_1(\mathbf{x}) + \alpha_2 f_2(\mathbf{x}) \text{ convex}$

$$f_1, \dots, f_n$$
 convex $\implies \max_i f_i(\mathbf{x})$ convex

- $f_1, \dots f_n$ convex $\implies \max_i f_i(\mathbf{x})$ - (Jensen's): $f(\mathbb{E}[\mathbf{X}]) \leq \mathbb{E}[f(\mathbf{X})]$
- Convex optimisation: minimise some convex function subject to some constraints: Minimise (over \mathbf{x})) $f_0(\mathbf{x})$ (where f_0 convex) such that $f_0(\mathbf{x})$ subject to $f_i(\mathbf{x}) \leq 0$ (where f_i convex) and $h_i(\mathbf{x}) = 0$ (where h_i affine (i.e. linear))
- Lagrangian: L(x, λ, ν) = f₀(x) + ∑_i λ_if_i(x) + ∑_i ν_ih_i(x) (λ and ν are called "dual variables"; can think of it like λ and ν are penalties that we want to minimise) min max_λ, ν, λ≥0</sub> L(x, λ, ν) = max_λ min L(x, λ, ν) due to nice

original problem equivalent dual problem convexity properties satisfying Minimax theorem; the RHS is often easier to solve (note: $\min_{\mathbf{x}} L(\mathbf{x}, \boldsymbol{\lambda}, \boldsymbol{\nu})$ is usually called $g(\boldsymbol{\lambda}, \boldsymbol{\nu})$) (if the min point of \mathbf{x} has $\lambda_i = 0$, then f_i is <u>inactive</u>)

• Karush-Kuhn-Tucker (KKT) Conditions (that the minimum point $(\mathbf{x}^*, \boldsymbol{\lambda}^*, \boldsymbol{\nu}^*)$ must satisfy):

- <u>Primal feasibility</u>: $f_i(\mathbf{x}^*) \le 0 \ \forall i = 1, \dots, m_{\text{ineq}}$ and $h_i(\mathbf{x}^*) = 0 \ \forall i = 1, \dots, m_{\text{eq}}$

- Dual feasibility: $\lambda_i^* \ge 0 \ \forall i = 1, \dots, m_{\text{ineq}}$

- <u>Complementary slackness</u>: $\lambda_i^* f_i(\mathbf{x}^*) = 0 \ \forall i = 1, ..., m_{\text{ineq}}$ (intuitively, either the min. pt. is on the boundary $(f_i(\mathbf{x}) = 0)$ or the constraint is inactive $(\lambda_i^* = 0)$)
- Vanishing gradient:

 $\overline{\nabla f_0(\mathbf{x}^*)} + \sum_{i=1}^{m_{\text{ineq}}} \lambda_i^* \nabla f_i(\mathbf{x}^*) + \sum_{i=1}^{m_{\text{eq}}} \nu_i^* \nabla h_i(\mathbf{x}^*) = \mathbf{0}$ In the general case this is necessary but not sufficient, but in the convex case it is both necessary and sufficient

• To solve a convex optimisation problem:

- write out the Lagrangian

- solve the dual problem by finding $g(\lambda, \nu)$, and then taking the maximum over λ, ν

• E.g. Hard margin SVM: minimise (over θ , θ_0) $\frac{1}{2} \|\theta\|^2$ subject to $\forall t$, $y_t \left(\theta^T \mathbf{x}_t + \theta_0\right) \ge 1$ - $L(\theta, \theta_0, \lambda) = \frac{1}{2} \|\theta\|^2 + \sum_{t=1}^n \lambda_t (1 - y_t(\theta^T \mathbf{x}_t + \theta_0))$ - $\frac{\partial L}{\partial \theta_0} = \sum_{t=1}^n (-\lambda_t y_t) = 0$ (equate to zero because we want the min point) - $\frac{\partial L}{\partial \theta} = \theta - \sum_{t=1}^n (\lambda_t y_t \mathbf{x}_t) = 0$ (equate to zero because we want the min point), so $\theta^* = \sum_{t=1}^n (\lambda_t y_t \mathbf{x}_t)$ - observe that at the min point, $\sum_{t=1}^n \lambda_t (1 - y_t(\theta^T \mathbf{x}_t + \theta_0))$ = $\sum_{t=1}^n \lambda_t - \sum_{t=1}^n \lambda_t y_t \theta^T \mathbf{x}_t - \sum_{t=1}^n \lambda_t y_t \theta_0$ = $\sum_{t=1}^n \lambda_t - \sum_{t=1}^n \lambda_t y_t (\sum_{t=1}^n (\lambda_t y_t \mathbf{x}_t))^T \mathbf{x}_t - \theta_0 \sum_{t=1}^n \lambda_t y_t$ = $\sum_{t=1}^n \lambda_t - \left(\sum_{s=1}^n (\lambda_s y_s \mathbf{x}_s)\right)^T \sum_{t=1}^n \lambda_t y_t \mathbf{x}_t$ = $\sum_{t=1}^n \lambda_t - \left\|\theta^*\|^2$ - so at the min point, $L(\theta, \theta_0, \lambda) = \sum_{t=1}^n \lambda_t - \frac{1}{2} \|\theta\|^2 = \sum_{t=1}^n \lambda_t - \frac{1}{2} \sum_{s=1}^n \sum_{t=1}^n \lambda_s \lambda_t y_s y_t \mathbf{x}_s^T \mathbf{x}_t$ - more formally, $g(\lambda) = \left\{\sum_{t=1}^{n} \lambda_t - \frac{1}{2} \sum_{s=1}^n \sum_{t=1}^{n} \lambda_s \lambda_t y_s y_t \mathbf{x}_s^T \mathbf{x}_t \text{ if } \sum_{t=1}^n \lambda_t y_t = 0 \right\}$ otherwise (since if $\sum_{t=1}^n \lambda_t y_t \neq 0$ we can choose θ_0 to be very big or very small, in order to make $\theta_0 \sum_{t=1}^n \lambda_t y_t \rightarrow -\infty$) - to find the maximum of $g(\lambda)$ (over $\lambda \ge 0$), solve maximise_{α} $\sum_{t=1}^n \alpha_t - \frac{1}{2} \sum_{s=1}^n \sum_{t=1}^n \alpha_s \alpha_t y_s y_t \mathbf{x}_s^T \mathbf{x}_t$ subject to $\alpha_t \ge 0$ and $\sum_t \alpha_t y_t = 0$

- then the classifier is $\boldsymbol{\theta} = \sum_{t=1}^{n} \alpha_t y_t \mathbf{x}_t$, and $\theta_0 = \frac{1}{y_t} - \boldsymbol{\theta}^T \mathbf{x}_t$ for any point (\mathbf{x}_t, y_t) where $\alpha_t > 0$ (i.e. $y_t(\boldsymbol{\theta}^T \mathbf{x}_t + \theta_0) = 1$, i.e. \mathbf{x}_t is a support vector)

• Computation of Primal vs Dual SVM:

Primal: Solving an optimisation problem with d variables and \overline{n} (or 2n) constraints (better for $n\gg d)$

<u>Dual</u>: Form $\frac{n^2}{2}$ kernel values (each of the \mathbf{x}_s , \mathbf{x}_t pairs), then solve an optimisation problem with n variables and n + 1constraints (better for $d \gg n$ if can compute $k(\mathbf{x}_s, \mathbf{x}_t)$ efficiently)

Boosting

• Decision stump: A linear classifier where the classifier must be either horizontal or vertical (with offset) (no diagonal lines)

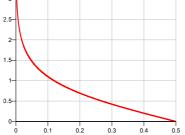
 $h(\mathbf{x}; \mathbf{\theta}) \coloneqq \operatorname{sgn}(s(x_k - \theta_0))$ where $\mathbf{\theta} = (s, k, \theta_0)$, where $s \in \{-1, 1\}, k \in \{1, \dots, d\}, \theta_0 \in \mathbb{R}$

(intuitively, s is the direction, k is the dimension index, θ_0 is the threshold/offset)

- we want to determine good $\boldsymbol{\theta}$

• Combined classifier: $\hat{y} \coloneqq \operatorname{sgn} \left(\sum_{m=1}^{M} \alpha_m h(\mathbf{x}; \boldsymbol{\theta}_m) \right)$ where $\alpha_m \ge 0$ are weights (i.e. the vote multiplier of base learner m) - we want to determine good $\boldsymbol{\theta}_1, \ldots, \boldsymbol{\theta}_M$

• Adaboost algorithm for learning $\theta_1, \ldots, \theta_M$ and $\alpha_1, \ldots, \alpha_M$ - Input: $\mathcal{D} = \{(\mathbf{x}_t, y_t)\}_{t=1}^n$, number of iterations M1. Init weights $w_0(t) \coloneqq \frac{1}{n}$ for $t = 1, \ldots, n$ 2. For $m = 1, \ldots, M$: a. Choose $h(\cdot; \hat{\theta}_m)$ as $\hat{\theta}_m \coloneqq \arg\min_{\theta} \sum_{y_t \neq h(\mathbf{x}_t; \theta)}^{n} w_{m-1}(t)$ (i.e. minimise the weight of all misclassified points) (note: $\arg\min_{\theta} \sum_{y_t \neq h(\mathbf{x}_t; \theta)}^{n} w_{m-1}(t) =$ $\arg\min_{\theta} \sum_{t=1}^n (-y_t h(\mathbf{x}_t; \theta)) w_{m-1}(t)$ since $-y_t h(\mathbf{x}_t; \theta) = 2 \cdot \mathbf{1}\{y_t \neq h(\mathbf{x}_t; \theta)\} - 1\}$ b. $\hat{\alpha}_m \coloneqq \frac{1}{2} \log \frac{1-\hat{\varepsilon}_m}{\hat{\varepsilon}_m}$ where $\hat{\varepsilon}_m \coloneqq \sum_{y_t \neq h(\mathbf{x}_t; \hat{\theta}_m)}^{n} w_{m-1}(t)$



(plot of $\hat{\alpha}_m$ against $\hat{\varepsilon}_m$:

c. Update weights: $w_m(t) \coloneqq \frac{1}{Z_m} w_{m-1}(t) e^{-y_t h(\mathbf{x}_t; \hat{\boldsymbol{\theta}}_m) \hat{\alpha}_m}$ where $Z_m \coloneqq \sum_{t=1}^n w_{m-1}(t) e^{-y_t h(\mathbf{x}_t; \hat{\boldsymbol{\theta}}_m) \hat{\alpha}_m}$ is the normalization factor (i.e. we want $\sum_{t=1}^n w_m(t) = 1$) (note: $e^{\hat{\alpha}_m} > 1$ and $e^{-\hat{\alpha}_m} < 1$ since $\hat{\alpha}_m > 0$, so it increases relative weights for misclassified points w.r.t. current decision stump)

3. The output classifier is $f_M(\mathbf{x}) \coloneqq \sum_{m=1}^M \hat{\alpha}_m h(\mathbf{x}; \hat{\boldsymbol{\theta}}_m)$

- Note: Adaboost never overfits points for some reason
- Training error: fraction of misclassified points: $\frac{1}{n} \sum_{t=1}^{n} \mathbf{1}\{y_t f_m(\mathbf{x}) \leq 0\}$

• **Theorem:** After M iterations, training error $\leq \exp\left(-2\sum_{m=1}^{M}\left(\frac{1}{2}-\hat{\varepsilon}_{m}\right)^{2}\right)$ (in particular, if $\hat{\varepsilon}_{m} \leq \frac{1}{2} - \gamma \ \forall m$, then training error $\leq \exp\left(-2M\gamma^{2}\right)$; and since training error must be a multiple of $\frac{1}{n}$, if $\exp\left(-2M\gamma^{2}\right) < \frac{1}{n}$ then training error must be zero (i.e. everything classified correctly))

• Adaboost proof:

1. Bound 0-1 loss with the exponential loss: $\frac{1}{n}\sum_{t=1}^{n} \mathbf{1}\{y_t f_M(\mathbf{x}) \leq 0\} \leq \frac{1}{n}\sum_{t=1}^{n} \exp(-y_t f_M(\mathbf{x})) \text{ (proof:} \\ \text{it is obvious)} \\ 2. By weight update formula,$ $<math display="block">w_M(t) = \frac{1}{n}\prod_{m=1}^{M} \frac{\exp(-y_t h(\mathbf{x}_t;\hat{\boldsymbol{\theta}}_m)\hat{\alpha}_m)}{Z_m} = \frac{1}{n}\frac{\exp(-\sum_{m=1}^{M} y_t h(\mathbf{x}_t;\hat{\boldsymbol{\theta}}_m)\hat{\alpha}_m)}{\prod_{m=1}^{M} Z_m} = \frac{1}{n}\frac{\exp(-y_t f_M(\mathbf{x}))}{\prod_{m=1}^{M} Z_m} \\ \text{Hence } 1 = \sum_{t=1}^{n} w_M(t) = \frac{1}{n}\frac{\sum_{t=1}^{n} \exp(-y_t f_M(\mathbf{x}))}{\prod_{m=1}^{M} Z_m}, \text{ and thus} \\ \frac{1}{n}\sum_{t=1}^{n} \exp(-y_t f_M(\mathbf{x})) = \prod_{m=1}^{M} Z_m \\ \text{Combining with step 1,} \\ \frac{1}{n}\sum_{t=1}^{n} \mathbf{1}\{y_t f_M(\mathbf{x}) \leq 0\} \leq \prod_{m=1}^{M} Z_m \\ 3. Z_m = \sum_{t=1}^{n} w_{m-1}(t)e^{-y_t h(\mathbf{x}_t;\hat{\boldsymbol{\theta}}_m)\hat{\boldsymbol{\alpha}}_m} = \\ \sum_{t:y_t \neq h(\mathbf{x}_t;\hat{\boldsymbol{\theta}}_m)} w_{m-1}(t)e^{\hat{\alpha}_m} + \sum_{t:y_t=h(\mathbf{x}_t;\hat{\boldsymbol{\theta}}_m)} w_{m-1}(t)e^{-\hat{\alpha}_m} \\ e^{\hat{\alpha}_m}\sum_{t:y_t \neq h(\mathbf{x}_t;\hat{\boldsymbol{\theta}}_m)} w_{m-1}(t) = \hat{\varepsilon}_m e^{\hat{\alpha}_m} + (1-\hat{\varepsilon}_m)e^{-\hat{\alpha}_m} \\ \text{So we want to pick } \hat{\alpha}_m \text{ to minimise } Z_m \\ (\text{proof skipped... the optimum } \hat{\alpha}_m = \frac{1}{2}\log\frac{1-\hat{\varepsilon}_m}{\hat{\varepsilon}_m}) \\ 4. \text{ Substitute } \hat{\alpha}_m \text{ to step 3, and get} \\ Z_m = \hat{\varepsilon}_m \sqrt{\frac{1-\hat{\varepsilon}_m}{\hat{\varepsilon}_m}} + (1-\hat{\varepsilon}_m)\sqrt{\frac{\hat{\varepsilon}_m}{1-\hat{\varepsilon}_m}} = 2\sqrt{\hat{\varepsilon}_m(1-\hat{\varepsilon}_m)} = \\ \end{array}$

$$\sqrt{1 - (1 - 2\hat{\varepsilon}_m)^2} = \exp\left(\frac{1}{2}\log\left(1 - (1 - 2\hat{\varepsilon}_m)^2\right)\right) \leq \exp\left(-\frac{1}{2}\left(1 - 2\hat{\varepsilon}_m\right)^2\right) = \exp\left(-2\left(\frac{1}{2} - \hat{\varepsilon}_m\right)^2\right)$$

(note: to remove log, use formula $\log(1 + a) \leq a \ \forall a \in \mathbb{R}$)
5. Substitute into step 2 end, we get
 $\frac{1}{n}\sum_{t=1}^n \mathbf{1}\{y_t f_M(\mathbf{x}) \leq 0\} \leq \exp\left(-2\sum_{m=1}^M \left(\frac{1}{2} - \hat{\varepsilon}_m\right)^2\right)$

• Claim: $\sum_{t:y_t \neq h(\mathbf{x}_t; \hat{\mathbf{\theta}}_m)} w_m(t) = \frac{1}{2}$ (hence the same $\hat{\mathbf{\theta}}$ will not be chosen twice in a row) <u>Proof</u>: LHS = $\frac{1}{2} \iff \sum_{t=1}^{n} (-y_t h(\mathbf{x}_t; \mathbf{\theta})) w_m(t) = 0$, and we have $\sum_{t=1}^{n} (-y_t h(\mathbf{x}_t; \mathbf{\theta})) w_m(t) =$ $\sum_{t:y_t=h(\mathbf{x}_t; \hat{\mathbf{\theta}}_m)} w_m(t) + \sum_{t:y_t \neq h(\mathbf{x}_t; \hat{\mathbf{\theta}}_m)} (-w_m(t)) =$ $\sum_{t:y_t=h(\mathbf{x}_t; \hat{\mathbf{\theta}}_m)} \frac{1}{Z_m} w_{m-1}(t) e^{-\hat{\alpha}_m} \sum_{t:y_t \neq h(\mathbf{x}_t; \hat{\mathbf{\theta}}_m)} \frac{1}{Z_m} w_{m-1}(t) e^{\hat{\alpha}_m} =$ $\frac{1}{Z_m} \left(e^{-\hat{\alpha}_m} (1 - \hat{\varepsilon}_m) - e^{\hat{\alpha}_m} \hat{\varepsilon}_m \right) = 0$ (last step is by definition of $\hat{\alpha}_m$ being chosen to minimise the expression obtained in step 3 of Adaboost proof)

Theory

Concentration

- Concentration: General idea: to show how well things concentrate around the mean:

 ℙ[|Y − m| > t] ≤ TailBound(t)
- Consider $Y_n = \frac{1}{n} \sum_{i=1}^n X_i$ where $\mathbb{E}[X_i] = \mu$ and $\operatorname{Var}[X_i] = \sigma^2$ and X_i are i.i.d. <u>Law of large numbers</u>: $\mathbb{P}[|Y_n - \mu| > \varepsilon] \to 0$ as $n \to \infty$ for any $\varepsilon > 0$ <u>Central limit theorem</u>: $\mathbb{P}[|Y_n - \mu| > \frac{\alpha}{\sqrt{n}}] \to 2\Phi(-\frac{\alpha}{\sigma})$ as $n \to \infty$ where Φ is the standard normal c.d.f. (inaccurate for very small probabilities)

Large deviations theory (important): $\mathbb{P}[|Y_n - \mu| > \varepsilon] < e^{-n \cdot \psi(\varepsilon)}$ (this is true for any ε and any n,

not just for large n)

• Basic inequalities:

 $\begin{array}{l} \frac{\text{Markov's ineq.: If } Z \text{ is a nonnegative rand. var., then} \\ \overline{\mathbb{P}[Z \geq t] \leq \frac{\mathbb{E}[Z]}{t}} \\ \frac{\text{Markov's ineq. for functions: If } Z \text{ is a rand. var. and } \phi \text{ is a nonnegative increasing function, then} \\ \overline{\mathbb{P}[Z \geq t]} \leq \overline{\mathbb{P}[\phi(Z) \geq \phi(t)]} \leq \frac{\mathbb{E}[\phi(Z)]}{\phi(t)} \\ \frac{\text{Chebyshev's ineq.: If } Z \text{ is a rand. var., then} \\ \overline{\mathbb{P}[|Z - \mathbb{E}[Z]| \geq t]} \leq \frac{\text{Var}[Z]}{t^2} \text{ (proof by letting } \phi(t) = t^2 \text{ and} \\ \text{replacing } Z \text{ by } |Z - \mathbb{E}[Z]|) \\ \frac{\text{Chernoff bound: If } Z \text{ is a rand. var. and } \lambda \geq 0, \text{ then} \\ \overline{\mathbb{P}[Z \geq t]} \leq e^{-\lambda t} \mathbb{E}[e^{\lambda Z}] \text{ (proof by letting } \phi(t) = e^{\lambda t}) \end{array}$

• Sum of independent random variables:

 $Z = X_1 + \dots + X_n \text{ and } Y_n = \frac{1}{n}Z \text{ where } X_i \text{ are i.i.d.:}$ $\frac{\text{Chebyshev's ineq.: } \mathbb{P}[|Y_n - \mathbb{E}[Y_n]| \ge \varepsilon] \le \frac{\text{Var}[X]}{n\varepsilon^2}$ $\frac{\text{Chernoff bound: } \mathbb{P}[Z \ge n\varepsilon] \le \exp(-n \cdot \psi_X^*(\varepsilon)) \text{ where } \psi_X^*(t) = \max_\lambda(\lambda t - \psi_X(\lambda)) \text{ and } \psi_X(\lambda) = \log \mathbb{E}[e^{\lambda X}]$ (e.g. for Gaussian $(X \sim N(0, \sigma^2)), \psi_X(\lambda) = \frac{\lambda^2 \sigma^2}{2}, \text{ so}$ $\psi_X^*(t) = \frac{t^2}{2\sigma^2}, \text{ so } \mathbb{P}[Z \ge n\varepsilon] \le \exp\left(-\frac{n\varepsilon^2}{2\sigma^2}\right), \text{ so}$

 $\mathbb{P}[|Z| \ge n\varepsilon] \le 2 \exp\left(-\frac{n\varepsilon^2}{2\sigma^2}\right))$ (note that for σ^2 -sub-Gaussian, we instead assume that $\psi_X(\lambda) \le \frac{\lambda^2 \sigma^2}{2}$, and the same result holds)

• Hoeffding's ineq.: $Z = X_1 + \dots + X_n$ where X_i are i.i.d. and $X_i \in [a_i, b_i]$: $\mathbb{P}\left[\frac{1}{n} | Z - \mathbb{E}[Z] | \ge \varepsilon\right] \le 2 \exp\left(\frac{-2n\varepsilon^2}{1 \sum_{i=1}^n (b_i - a_i)^2}\right)$ If $a_i = a$ and $b_i = b$ for all *i* then: $\mathbb{P}\left[\frac{1}{n}|Z - \mathbb{E}[Z]| \ge \varepsilon\right] \le 2\exp\left(\frac{-2n\varepsilon^2}{(b-a)^2}\right)$ If $a_i = 0$ and $b_i = 1$ for all *i* then: $\mathbb{P}\left[\frac{1}{n}|Z - \mathbb{E}[Z]| \ge \varepsilon\right] \le 2\exp\left(-2n\varepsilon^2\right)$

Statistical learning theory

• **Underfitting**: High training error (and hence high test error)

 $\ensuremath{\textbf{Overfitting:}}$ Low training error but high test error

• Setup:

- Data drawn from distribution $P_{\mathbf{X}Y}$ (unknown), i.e. $\mathcal{D} = \{(\mathbf{x}_t, y_t)\}_{t=1}^n$, where each $(\mathbf{x}_t, y_t) \sim P_{\mathbf{X}Y}$ and i.i.d. - Loss function $\ell(y, \hat{y})$ where $\hat{y} = f(x)$ (where f is our classifier) (ℓ could be any loss function) - Training error (empirical risk): $\overline{R}_n(f) = \frac{1}{n} \sum_{t=1}^n \ell(y_t, f(\mathbf{x}_t))$ - Test error (true risk): $R(f) = \mathbb{E}[\ell(y, f(\mathbf{x}))]$ where $(x,y) \sim P_{\mathbf{X}Y}$ $-\underbrace{R(f)}_{\text{Test error}} = \underbrace{R_n f}_{\text{Training error}} + \underbrace{(R(f) - R_n f)}_{\text{Generalisation error}}$ (Generalisation error is large if we overfit) - Consider algorithm that outputs $f \in \mathcal{F}$ (if \mathcal{F} is too small then we underfit; if \mathcal{F} is too large then we overfit) • Task: Let $f_{\text{erm}} = \arg\min_{f \in \mathcal{F}} R_n(f)$, want to find out if $f_{\rm erm}$ has small R(f) too? Will show that $R(f_{\rm erm}) \leq R(f^*) + \varepsilon$ with probability $\geq 1 - \delta$ provided that $n \geq \overline{n}(\varepsilon, \delta)$ (where $f^* = \arg\min_{f \in \mathcal{F}} R(f)$, and \overline{n} is known as the sample complexity) (called the "probably approximately correct" (PAC) guarantee) (the analysis is true even for the worst case $P_{\mathbf{X}Y}$, so for good distributions, it may be possible to use less data points) \mathcal{F} is PAC-learnable iff this is attainable for all $\varepsilon > 0, \delta > 0$ with $\overline{n} < \infty$ (regardless of $P_{\mathbf{X}Y}$) (when $|\mathcal{F}|$ is small enough this is usually possible) Realisable: $y_i = f(\mathbf{x}_i)$ for all *i*, for some $f \in \mathcal{F}$ Agnostic: Just compare with the (ideal) $f^* \in \mathcal{F}$ (we usually use this) • PAC guarantee for finite \mathcal{F} : If $|\mathcal{F}| < \infty$ and $\ell(y, \hat{y}) \in [0, 1]$, then \mathcal{F} is PAC-learnable (i.e. $R(f_{\rm erm}) \leq R(f^*) + \varepsilon$ with probability $\geq 1 - \delta$) with $\overline{n}_{\mathcal{F}}(\varepsilon,\delta) = \frac{2}{c^2} \log \frac{2|\mathcal{F}|}{\delta}$

(equivalently, $R(f_{\text{erm}}) - R(f^*) \le \varepsilon \le \sqrt{\frac{2}{n} \log \frac{2|\mathcal{F}|}{\delta}}$)

- Proof of PAC guarantee for finite \mathcal{F} : 1. Fix $f \in \mathcal{F}$. Let $z_i = \ell(y_i, f(\mathbf{x}_i)) \forall i$. Then since (\mathbf{x}_i, y_i) are i.i.d., z_i is also i.i.d.. Since $z_i \in [0, 1]$, and $\mathbb{E}[z_i] = \mathbb{E}[\ell(y_i, f(\mathbf{x}_i))] = R(f)$, we apply Hoeffding's ineq to get $\mathbb{P}\left[\left|\frac{1}{n}\sum_{i=1}^{n} z_i - R(f)\right| > \varepsilon_0\right] \le 2 \exp\left(-2n\varepsilon_0^2\right)$. Then observe that $\frac{1}{n}\sum_{i=1}^{n} z_i = \frac{1}{n}\sum_{i=1}^{n}\ell(y_i, f(\mathbf{x}_i)) = R_n(f)$, so $\mathbb{P}\left[\left|R_n(f) - R(f)\right| > \varepsilon_0\right] \le 2 \exp\left(-2n\varepsilon_0^2\right)$. 2. Take the union bound: $\mathbb{P}\left[\bigcup_{f \in \mathcal{F}}\left\{\left|R_n(f) - R(f)\right| > \varepsilon_0\right\}\right] \le 2\left|\mathcal{F}\right|\exp\left(-2n\varepsilon_0^2\right)$ 3. So we pick $\delta = 2\left|\mathcal{F}\right|\exp\left(-2n\varepsilon_0^2\right)$, so $n = \frac{1}{2\varepsilon_0^2}\log\frac{2|\mathcal{F}|}{\delta}$. 4. Suppose $|R_n(f) - R(f)| \le \varepsilon_0 \ \forall f \in \mathcal{F}$ (holds with probability $\ge 1 - \delta$), then: $R(f_{\text{erm}}) - R(f^*) = \frac{R(f_{\text{erm}}) - R_n(f_{\text{erm}}) + \frac{R_n(f_{\text{erm}}) - R_n(f^*) + R_n(f^*) - R(f^*)}{\le \varepsilon_0}}{\le 2\varepsilon_0}$.
- Shattering: A set of points $\mathbf{x}_1, \ldots, \mathbf{x}_k$ is said to be shattered by \mathcal{F} if $|\{(f(\mathbf{x}_1), \ldots, f(\mathbf{x}_k)) \mid f \in \mathcal{F}\}| = 2^k$

• VC Dimension: $d_{VC} = d_{VC}(\mathcal{F}) = \text{largest } k$ such that $\exists \mathbf{x}_1, \ldots, \mathbf{x}_k$ that \mathcal{F} shatters (i.e. largest k such that $\exists \mathbf{x}_1, \ldots, \mathbf{x}_k$ for which all combinations of labels (i.e. $f(\mathbf{x}_1), \ldots, f(\mathbf{x}_k)$) can be produced by classifiers in \mathcal{F}) (to determine the VC dimension, must show that k works but k + 1 does not work)

• Corollary for Finite \mathcal{F} : If \mathcal{F} is finite, $d_{\rm VC} \leq \log_2 |\mathcal{F}|$

• Sauer's lemma: Effective #f in $\mathcal{F} \leq \sum_{i=0}^{d_{\mathrm{VC}}} {n \choose i}$ Slightly weaker version: Effective #f in $\mathcal{F} \leq (n+1)^{d_{\mathrm{VC}}}$

• PAC guarantee for infinite \mathcal{F} : Assuming 0-1 loss, \mathcal{F} is PAC-learnable with $\overline{n}_{\mathcal{F}}(\varepsilon, \delta) = C \cdot \frac{d_{\mathrm{VC}}(\mathcal{F}) + \log \frac{1}{\delta}}{\varepsilon^2}$

• Converse guarantee: If $d_{\rm VC} = \infty$, then \mathcal{F} is not PAC-learnable (though there may be certain choices of $P_{\mathbf{X}Y}$ where learning might still work)

• Examples

- Linear classifier (without offset): $d_{VC} = d$ (the dimension of the space) - Linear classifier (with offset): $d_{VC} = d + 1$

Unsupervised Learning

• Training data: $\mathcal{D} = {\mathbf{x}_t}_{t=1}^n$

• K-means clustering:

<u>Goal</u>: Partition \mathcal{D} into clusters $\mathcal{D}_1, \ldots, \mathcal{D}_k$ such that the associated cluster centres $\boldsymbol{\mu}_1, \ldots, \boldsymbol{\mu}_K \in \mathbb{R}^d$ minimise $J\left(\{\mathcal{D}_j\}_{j=1}^K, \{\boldsymbol{\mu}_j\}_{j=1}^K\right) = \sum_{j=1}^K \sum_{\mathbf{x} \in \mathcal{D}_j} \|\mathbf{x} - \boldsymbol{\mu}_j\|^2$ (i.e. minimise the sum of squared distance from the respective cluster centres)

• K-means algorithm:

1. Minimise $J\left(\{\mathcal{D}_j\}_{j=1}^K, \{\mathbf{\mu}_j\}_{j=1}^K\right)$ w.r.t. $\{\mathcal{D}_j\}_{j=1}^K$ for fixed $\{\mathbf{\mu}_j\}_{j=1}^K$ (i.e. associate each data point with the cluster centre closest to it)

2. Minimise $J\left(\{\mathcal{D}_j\}_{j=1}^K, \{\boldsymbol{\mu}_j\}_{j=1}^K\right)$ w.r.t. $\{\boldsymbol{\mu}_j\}_{j=1}^K$ for fixed $\{\mathcal{D}_j\}_{j=1}^K$ (i.e. set each cluster centre to the mean of the data points associated with it, i.e. $\boldsymbol{\mu}_j = \frac{1}{|\mathcal{D}_j|} \sum_{\mathbf{x} \in \mathcal{D}_j} \mathbf{x}$) 3. Repeat until no change

Notes:

- initially, choose $\{\mu_j\}_{j=1}^K$ at random

- K-means finds a local minimum for J

• Distribution learning:

Goal: Estimate a distribution $\hat{p}(\mathbf{x})$ that models the data well (either p.m.f. (discrete) or p.d.f. (continuous)) - assume $p(\mathbf{x})$ is determined by $\boldsymbol{\theta}$, i.e. $p(\mathbf{x}; \boldsymbol{\theta})$

- to find $\hat{\theta}$, we use maximum likelihood, i.e.

 $\hat{\mathbf{\theta}} = \arg \max_{\mathbf{\theta}} \prod_{t=1}^{n} p(\mathbf{x}_t; \mathbf{\theta}) = \arg \max_{\mathbf{\theta}} \sum_{t=1}^{n} \log p(\mathbf{x}_t; \mathbf{\theta})$